

DIGITAL SIGNAL PROCESSING

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To Phyllis and Dorothy

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$$c_{xx}(m) = \sum_{n=-(N-1)}^{N-1} c_{xx}(m) e^{-j\omega m} \quad (11.24)$$

of the real finite-length sequence $x(n)$, $0 \leq n \leq N-1$.

$$X(e^{j\omega}) = \sum_{n=0}^{N-1} x(n) e^{-j\omega n}$$

Problem 1 of this chapter).

$$I_N(\omega) = \frac{1}{N} |X(e^{j\omega})|^2 \quad (11.25)$$

$I_N(\omega)$ is often called the *periodogram*.

It is best to determine the bias and variance of the estimate of the power spectrum. The expected value of

$$E[I_N(\omega)] = \sum_{m=-(N-1)}^{N-1} E[c_{xx}(m)] e^{-j\omega m} \quad (11.26)$$

for a zero mean process

$$\phi_{xx}(m) = \frac{N - |m|}{N} \phi_{xx}(m), \quad |m| < N$$

$$E[c_{xx}(m)] = \sum_{n=-(N-1)}^{N-1} \left(\frac{N - |m|}{N} \right) \phi_{xx}(m) e^{-j\omega m} \quad (11.27)$$

limits of summation and the factor $(N - |m|)/N$, is the Fourier transform of $\phi_{xx}(m)$, and therefore the expected value of the power spectrum, $P_{xx}(\omega)$.

$E[I_N(\omega)]$ is the Fourier transform of the estimate $c'_{xx}(m)$; i.e.,

$$E[I_N(\omega)] = \sum_{m=-(N-1)}^{N-1} c'_{xx}(m) e^{-j\omega m} \quad (11.28)$$

is

$$E[I_N(\omega)] = \sum_{m=-(N-1)}^{N-1} E[c'_{xx}(m)] e^{-j\omega m} \quad (11.29)$$

$$= \sum_{m=-(N-1)}^{N-1} \phi_{xx}(m) e^{-j\omega m}$$

windowed autocorrelation sequences. In the case of Eq. (11.27) the window is the triangular window

$$w_B(m) = \begin{cases} \frac{N - |m|}{N}, & |m| < N \\ 0, & \text{otherwise} \end{cases} \quad (11.28)$$

In Chapter 5 we called this the Bartlett window. For Eq. (11.29) the window is rectangular; i.e.,

$$w_R(n) = \begin{cases} 1, & |m| < N \\ 0, & \text{otherwise} \end{cases} \quad (11.29)$$

Using the concepts introduced in Chapter 5 we can see that Eqs. (11.28) and (11.29) can be interpreted in the frequency domain as the convolution

$$E[I_N(\omega)] = \frac{1}{2\pi} \int_{-\pi}^{\pi} P_{xx}(\theta) W_B(e^{j(\omega-\theta)}) d\theta \quad (11.30)$$

and

$$E[P_N(\omega)] = \frac{1}{2\pi} \int_{-\pi}^{\pi} P_{xx}(\theta) W_R(e^{j(\omega-\theta)}) d\theta \quad (11.31)$$

where

$$W_B(e^{j\omega}) = \frac{1}{N} \left(\frac{\sin [\omega N/2]}{\sin [\omega/2]} \right)^2 \quad (11.32)$$

and

$$W_R(e^{j\omega}) = \frac{\sin [\omega(2N - 1)/2]}{\sin [\omega/2]} \quad (11.33)$$

are the Fourier transforms of the Bartlett and rectangular windows, respectively.

11.3.2 Variance of the Periodogram

To obtain an expression for the variance of the periodogram, it is convenient to first assume that the sequence $x(n)$, $0 \leq n \leq N-1$, is a sample of a real, white, zero-mean process with Gaussian probability density functions. The periodogram $I_N(\omega)$ can be expressed as

$$I_N(\omega) = \frac{1}{N} |X(e^{j\omega})|^2$$

$$= \frac{1}{N} \sum_{l=0}^{N-1} \sum_{m=0}^{N-1} x(l)x(m) e^{j\omega m} e^{-j\omega l}$$

11.3.1 Definition of the Periodogram

As an estimate of the power density spectrum let us consider the Fourier transform of the biased autocorrelation estimate $c_{xx}(m)$. That is,

$$I_N(\omega) = \sum_{m=-(N-1)}^{N-1} c_{xx}(m)e^{-j\omega m} \quad (11.24)$$

Since the Fourier transform of the real finite-length sequence $x(n)$, $0 \leq n \leq N-1$, is

$$X(e^{j\omega}) = \sum_{n=0}^{N-1} x(n)e^{-j\omega n}$$

it can be shown that (see Problem 1 of this chapter).

$$I_N(\omega) = \frac{1}{N} |X(e^{j\omega})|^2 \quad (11.25)$$

The spectrum estimate $I_N(\omega)$ is often called the *periodogram*.

As before, it is of interest to determine the bias and variance of the periodogram as an estimate of the power spectrum. The expected value of $I_N(\omega)$ is

$$E[I_N(\omega)] = \sum_{m=-(N-1)}^{N-1} E[c_{xx}(m)]e^{-j\omega m} \quad (11.26)$$

Since we have shown that for a zero mean process

$$E[c_{xx}(m)] = \frac{N - |m|}{N} \phi_{xx}(m), \quad |m| < N$$

then

$$E[I_N(\omega)] = \sum_{m=-(N-1)}^{N-1} \left(\frac{N - |m|}{N} \right) \phi_{xx}(m)e^{-j\omega m} \quad (11.27)$$

Thus because of the finite limits of summation and the factor $(N - |m|)/N$, $E[I_N(\omega)]$ is not equal to the Fourier transform of $\phi_{xx}(m)$, and therefore the periodogram is a biased estimate of the power spectrum, $P_{xx}(\omega)$.

Alternatively, consider the Fourier transform of the estimate $c'_{xx}(m)$; i.e.,

$$P_N(\omega) = \sum_{m=-(N-1)}^{N-1} c'_{xx}(m)e^{-j\omega m} \quad (11.28)$$

The expected value of $P_N(\omega)$ is

$$E[P_N(\omega)] = \sum_{m=-(N-1)}^{N-1} E[c'_{xx}(m)]e^{-j\omega m} \quad (11.29)$$

Again, because of the finite limits of summation $P_N(\omega)$, even though $c'_{xx}(m)$ is an unbiased estimator of $\phi_{xx}(m)$,

We can interpret Eqs. (11.27) and (11.29) as the Fourier transforms of windowed autocorrelation sequences. In the case of $I_N(\omega)$ the window is the triangular window

$$w_B(m) = \begin{cases} \frac{N - |m|}{N}, & |m| < N \\ 0, & \text{otherwise} \end{cases}$$

In Chapter 5 we called this the Bartlett window. It is rectangular; i.e.,

$$w_R(n) = \begin{cases} 1, & |m| < N \\ 0, & \text{otherwise} \end{cases}$$

Using the concepts introduced in Chapter 5, Eqs. (11.27) and (11.29) can be interpreted in the frequency domain as

$$E[I_N(\omega)] = \frac{1}{2\pi} \int_{-\pi}^{\pi} P_{xx}(\theta)W_B(\omega - \theta)d\theta$$

and

$$E[P_N(\omega)] = \frac{1}{2\pi} \int_{-\pi}^{\pi} P_{xx}(\theta)W_R(\omega - \theta)d\theta$$

where

$$W_B(e^{j\omega}) = \frac{1}{N} \left(\frac{\sin [\omega N/2]}{\sin [\omega/2]} \right)^2$$

and

$$W_R(e^{j\omega}) = \frac{\sin [\omega(2N - 1)/2]}{\sin [\omega/2]}$$

are the Fourier transforms of the Bartlett and rectangular windows, respectively.

11.3.2 Variance of the Periodogram

To obtain an expression for the variance of the periodogram it is convenient to first assume that the sequence $x(n)$, $0 \leq n \leq N-1$, is a real, white, zero-mean process with Gaussian random functions. The periodogram $I_N(\omega)$ can be expressed as

$$I_N(\omega) = \frac{1}{N} |X(e^{j\omega})|^2$$

To evaluate the covariance of $I_N(\omega)$ at two frequencies ω_1 and ω_2 we first consider

$$E[I_N(\omega_1)I_N(\omega_2)] = \frac{1}{N^2} \sum_{k=0}^{N-1} \sum_{l=0}^{N-1} \sum_{m=0}^{N-1} \sum_{n=0}^{N-1} E[x(k)x(l)x(m)x(n)] e^{j[\omega_1(k-l) + \omega_2(m-n)]} \quad (11.36)$$

To obtain a useful result, we must simplify Eq. (11.36). In general, it is not possible to obtain a very simple result even when $x(n)$ is white, because $E[x(n)x(n+m)] = \sigma_x^2 \delta(m)$ does not guarantee a simple expression for $E[x(k)x(l)x(m)x(n)]$ for all combinations of $k, l, m,$ and n . However, in the case of a white Gaussian process, it can be shown [7] that

$$\begin{aligned} E[x(k)x(l)x(m)x(n)] &= E[x(k)x(l)]E[x(m)x(n)] \\ &\quad + E[x(k)x(m)]E[x(l)x(n)] \\ &\quad + E[x(k)x(n)]E[x(l)x(m)] \end{aligned}$$

Therefore,

$$E[x(k)x(l)x(m)x(n)] = \begin{cases} \sigma_x^4, & k = l \text{ and } m = n \\ & \text{or } k = m \text{ and } l = n \\ & \text{or } k = n \text{ and } l = m \\ 0, & \text{otherwise} \end{cases} \quad (11.37)$$

For other than Gaussian joint density functions, the result will not necessarily be so simple. However, our objective is to give a result that will lend insight into the problems of spectrum estimation rather than to give a general formula with wide validity which would be difficult to interpret. Thus, if we substitute Eq. (11.37) into Eq. (11.36), we obtain

$$E[I_N(\omega_1)I_N(\omega_2)] = \frac{\sigma_x^4}{N^2} \left\{ N^2 + \sum_{m=0}^{N-1} \sum_{n=0}^{N-1} e^{j(m-n)(\omega_1+\omega_2)} + \sum_{m=0}^{N-1} \sum_{n=0}^{N-1} e^{j(n-m)(\omega_1-\omega_2)} \right\}$$

or

$$\begin{aligned} E[I_N(\omega_1)I_N(\omega_2)] &= \sigma_x^4 \left\{ 1 + \left(\frac{\sin [(\omega_1 + \omega_2)N/2]}{N \sin [(\omega_1 + \omega_2)/2]} \right)^2 \right. \\ &\quad \left. + \left(\frac{\sin [(\omega_1 - \omega_2)N/2]}{N \sin [(\omega_1 - \omega_2)/2]} \right)^2 \right\} \end{aligned} \quad (11.38)$$

(If the signal is non-Gaussian, Eq. (11.38) contains additional terms which are proportional to $1/N$ [4, 8].) The covariance of the periodogram is

$$\text{cov} [I_N(\omega_1), I_N(\omega_2)] = E[I_N(\omega_1)I_N(\omega_2)] - E[I_N(\omega_1)]E[I_N(\omega_2)] \quad (11.39)$$

Since $E[I_N(\omega_1)] = E[I_N(\omega_2)] =$

$\text{cov} [I_N(\omega_1), I_N(\omega_2)] =$

From Eq. (11.40) we can draw the periodogram. The variance of the periodogram at frequency $\omega = \omega_1 = \omega_2$ is

$$\text{var} [I_N(\omega)] = \text{cov} [I_N(\omega), I_N(\omega)]$$

Clearly, the variance of $I_N(\omega)$ goes to infinity. Thus the periodogram is not a consistent estimator. The variance of $I_N(\omega)$ is of the order of σ_x^4/N .

We also see from Eq. (11.40) that the variance of $I_N(\omega)$ is of the order of $2\pi l/N$, where k and l are integers.

$$\text{cov} [I_N(\omega_1), I_N(\omega_2)] = \sigma_x^4 \left\{ \left(\frac{\sin [(\omega_1 + \omega_2)N/2]}{N \sin [(\omega_1 + \omega_2)/2]} \right)^2 + \left(\frac{\sin [(\omega_1 - \omega_2)N/2]}{N \sin [(\omega_1 - \omega_2)/2]} \right)^2 \right\}$$

which is equal to zero for $k \neq l$. The periodogram is not a consistent estimator of the power spectrum by integer multiples of the frequency. These uncorrelated frequency samples are plotted together. It is reasonable to expect that the periodogram should approach a constant value if the signal was white. A consequence of this is that the periodogram approaches a non-zero value as the spectral samples with zero covariance. As the record length becomes longer, the variance of the periodogram increases. This behavior is illustrated in the periodogram plotted for record lengths of 100, 200, and 400 samples.

11.3.3 General Variance Expression

All the previous discussion has been for the case of white noise. If we consider a general signal, the analysis is considerably more difficult. In this more general case, we will approach and develop an approach to the problem. The approach we shall take is heuristic; a more rigorous approach is given by Watts [5]. With the application of the approximate results derived

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