EXHIBIT 4 PART



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overlap between the component densities, one can expect biased estimates and less than optimal results.

Despite these drawbacks, the simplicity of decision-directed procedures makes the Bayesian approach computationally feasible, and a flawed solution is often better than none. If conditions are favorable, performance that is nearly optimal can be achieved at far less computational expense. The literature contains a few rather complicated analyses of particular decision-directed procedures, and numerous reports of experimental results. The basic conclusions are that most of these procedures work well if the parametric assumptions are valid, if there is little overlap between the component densities, and if the initial classifier design is at least roughly correct.

6.6 DATA DESCRIPTION AND CLUSTERING

Let us reconsider our original problem of learning something of use from a set of unlabelled samples. Viewed geometrically, these samples form clouds of points in a d-dimensional space. Suppose that we knew that these points came from a single normal distribution. Then the most we could learn from the data would be contained in the sufficient statistics—the sample mean and the sample covariance matrix. In essence, these statistics constitute a compact description of the data. The sample mean locates the center of gravity of the cloud. It can be thought of as the single point x that best represents all of the data in the sense of minimizing the sum of squared distances from x to the samples. The sample covariance matrix tells us how well the sample mean describes the data in terms of the amount of scatter that exists in various directions. If the data points are actually normally distributed, then the cloud has a simple hyperellipsoidal shape, and the sample mean tends to fall in the region where the samples are most densely concentrated.

Of course, if the samples are not normally distributed, these statistics can give a very misleading description of the data. Figure 6.7 shows four different data sets that all have the same mean and covariance matrix. Obviously, second-order statistics are incapable of revealing all of the structure in an arbitrary set of data.

By assuming that the samples come from a mixture of c normal distributions, we can approximate a greater variety of situations. In essence, this corresponds to assuming that the samples fall in hyperellipsoidally-shaped clouds of various sizes and orientations. If the number of component densities is not limited, we can approximate virtually any density function in this way, and use the parameters of the mixture to describe the data. Unfortunately, we have seen that the problem of estimating the parameters of a mixture



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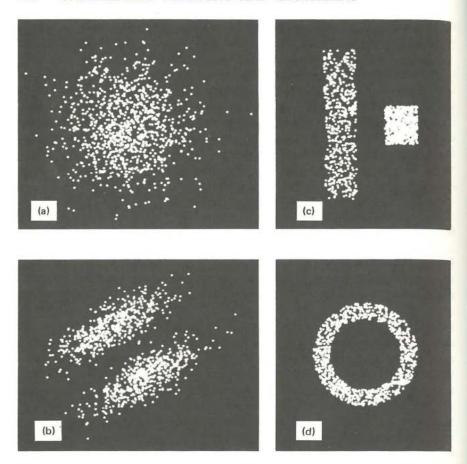


FIGURE 6.7. Data sets having identical second-order statistics.

density is not trivial. Furthermore, in situations where we have relatively little a priori knowledge about the nature of the data, the assumption of particular parametric forms may lead to poor or meaningless results. Instead of finding structure in the data, we would be imposing structure on it.

One alternative is to use one of the nonparametric methods described in Chapter 4 to estimate the unknown mixture density. If accurate, the resulting estimate is certainly a complete description of what we can learn from the data. Regions of high local density, which might correspond to significant subclasses in the population, can be found from the peaks or modes of the estimated density.



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If the goal is to find subclasses, a more direct alternative is to use a clustering procedure. Roughly speaking, clustering procedures yield a data description in terms of clusters or groups of data points that possess strong internal similarities. The more formal procedures use a criterion function, such as the sum of the squared distances from the cluster centers, and seek the grouping that extremizes the criterion function. Because even this can lead to unmanageable computational problems, other procedures have been proposed that are intuitively appealing but that lead to solutions having no established properties. Their use is usually justified on the ground that they are easy to apply and often yield interesting results that may guide the application of more rigorous procedures.

6.7 SIMILARITY MEASURES

Once we describe the clustering problem as one of finding natural groupings in a set of data, we are obliged to define what we mean by a natural grouping. In what sense are we to say that the samples in one cluster are more like one another than like samples in other clusters? This question actually involves two separate issues—how should one measure the similarity between samples, and how should one evaluate a partitioning of a set of samples into clusters? In this section we address the first of these issues.

The most obvious measure of the similarity (or dissimilarity) between two samples is the distance between them. One way to begin a clustering investigation is to define a suitable distance function and compute the matrix of distances between all pairs of samples. If distance is a good measure of dissimilarity, then one would expect the distance between samples in the same cluster to be significantly less than the distance between samples in different clusters.

Suppose for the moment that we say that two samples belong to the same cluster if the Euclidean distance between them is less than some threshold distance d_0 . It is immediately obvious that the choice of d_0 is very important. If d_0 is very large, all of the samples will be assigned to one cluster. If d_0 is very small, each sample will form an isolated cluster. To obtain "natural" clusters, d_0 will have to be greater than typical within-cluster distances and less than typical between-cluster distances (see Figure 6.8).

Less obvious perhaps is the fact that the results of clustering depend on the choice of Euclidean distance as a measure of dissimilarity. This choice implies that the feature space is isotropic. Consequently, clusters defined by Euclidean distance will be invariant to translations or rotations—rigid-body motions of the data points. However, they will not be invariant to linear transformations in general, or to other transformations that distort the



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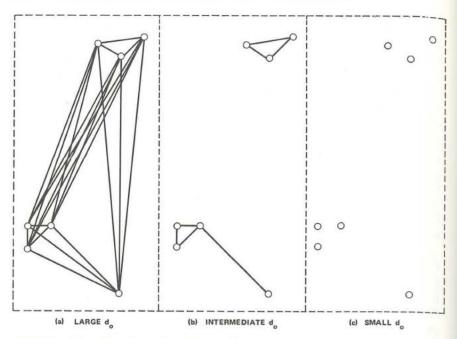


FIGURE 6.8. The effect of a distance threshold on clustering (Lines are drawn between points closer than a distance d_0 apart).

distance relationships. Thus, as Figure 6.9 illustrates, a simple scaling of the coordinate axes can result in a different grouping of the data into clusters. Of course, this is of no concern for problems in which arbitrary rescaling is an unnatural or meaningless transformation. However, if clusters are to mean anything, they should be invariant to transformations natural to the problem.

One way to achieve invariance is to normalize the data prior to clustering. For example, to obtain invariance to displacement and scale changes, one might translate and scale the axes so that all of the features have zero mean and unit variance. To obtain invariance to rotation, one might rotate the axes so that they coincide with the eigenvectors of the sample covariance matrix. This transformation to principal components can be preceded and/or followed by normalization for scale.

However, the reader should not conclude that this kind of normalization is necessarily desirable. Consider, for example, the matter of translating and scaling the axes so that each feature has zero mean and unit variance. The rationale usually given for this normalization is that it prevents certain features from dominating distance calculations merely because they have



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